

1    **Does extreme climate event exposure influence climate-related opinions? The**  
2    **case of the 2019–2020 Australian Black Summer bushfires**

3    Matthew Andreotta<sup>1</sup>, Fabio Boschetti<sup>1</sup>, Simon Farrell<sup>2</sup>, Cécile Paris<sup>3</sup>, Iain Walker<sup>4</sup>, and  
4    Mark J. Hurlstone<sup>5</sup>

5    <sup>1</sup>Environment, CSIRO, Australia

6    <sup>2</sup>School of Psychological Science, University of Western Australia, Australia

7    <sup>3</sup>Data61, CSIRO, Australia

8    <sup>4</sup>School of Psychological Sciences, University of Melbourne, Australia

9    <sup>5</sup>Department of Psychology, Lancaster University, UK

10    **Author Note**

11    Matthew Andreotta  <https://orcid.org/0000-0002-2996-6031>

12    Fabio Boschetti  <https://orcid.org/0000-0001-8999-6913>

13    Simon Farrell  <https://orcid.org/0000-0001-7452-8789>

14    Cécile Paris  <https://orcid.org/0000-0003-3816-0176>

15    Iain Walker  <https://orcid.org/0000-0002-1020-5873>

16    Mark J. Hurlstone  <https://orcid.org/0000-0001-9920-6284>

17    Correspondence concerning this article should be addressed to Mark Hurlstone,

18    Department of Psychology, Lancaster University, Lancaster, United Kingdom, LA1 4YW.

19    email: [m.hurlstone@lancaster.ac.uk](mailto:m.hurlstone@lancaster.ac.uk)

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21    <https://github.com/matt-lab/bushfire-audience-segmentation>.

22        **Code availability statement.** Analysis scripts are available online, at:

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25    Conceptualisation, Data curation, Formal analysis, Investigation, Methodology,

26    Visualisation, Writing—original draft, Writing—review and editing. **Fabio Boschetti**,

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37

## Abstract

38 We report three studies examining the impact of the 2019-2020 Australian bushfires, known  
39 as the Black Summer, on Australians' climate-related opinions. Study 1 was conducted  
40 before the peak of the bushfires, whereas Studies 2 and 3 were conducted afterwards. In all  
41 studies, respondents completed a Q-sort task ranking opinion statements about climate  
42 change by degree of endorsement. Study 3 also incorporated measures of bushfire  
43 perceptions and climate policy support. Q-sort responses consistently revealed evidence for  
44 three opinion segments: climate-change Acceptors, Fencesitters, and Sceptics. Over time,  
45 the proportion of Acceptors decreased, the proportion of Fencesitters increased, while  
46 Sceptics remained stable. Perceptions of the bushfires varied across segments. Although all  
47 segments perceived the fires as severe, Acceptors tended to attribute them to climate  
48 change, whereas Fencesitters and Sceptics attributed them to mass arson. However, even  
49 many Acceptors endorsed the mass-arson claim. On climate policy, Acceptors favoured  
50 stronger action, Fencesitters were evenly divided, and Sceptics mostly opposed change. Our  
51 results suggest the Black Summer bushfires did not elicit greater acceptance of  
52 anthropogenic climate change or heightened concern. Instead, misinformation, particularly  
53 conservative media narratives attributing the fires to mass arson, may have influenced  
54 beliefs about the fires' causes, especially among undecided and sceptical individuals.

55

56 *Keywords:* Black Summer bushfires · climate change · climate opinion · personal  
57 experience · extreme events · event attribution · misinformation

58 **Does extreme climate event exposure influence climate-related opinions? The**  
59 **case of the 2019–2020 Australian Black Summer bushfires**

60 In October 2019, lightning ignited the largest fire in Australia's recent history (Rural  
61 Fire Service, 2020). The 'megafire', so-called for its intensity, size, and difficulty to control,  
62 endured for fifteen weeks, burning 512,000 hectares of land, including the Blue Mountains  
63 World Heritage Area (Rural Fire Service, 2020). Accompanying the megafire were bushfires  
64 in all Australian states and territories throughout the unprecedented 2019-2020 bushfire  
65 season (Boer et al., 2020), which became known as the Black Summer. Together, these fires  
66 directly killed at least 33 people, burnt over 24 million hectares, destroyed over 3,000  
67 homes, killed or displaced nearly three billion animals, and affected nearly 80% of  
68 Australians either directly or indirectly (Hughes et al., 2020; The Royal Commission into  
69 National Natural Disaster Arrangements, 2020). On a local scale, the fires trapped  
70 thousands of Australians without essential goods and services (The Royal Commission into  
71 National Natural Disaster Arrangements, 2020). On a national and international scale, the  
72 fires transformed the usually festive season into one of grief and vigilance (Head, 2020).

73 Given the extended time scale over which the bushfires raged, the harrowing and  
74 sustained media reporting of the devastation caused, and the large proportion of  
75 Australians affected by the events, a natural question to ask is whether this Black Summer  
76 of crisis altered Australians' climate-related opinions. More generally, the Black Summer  
77 presents an opportunity to examine how extreme events of this magnitude may shape  
78 public opinions about climate change. In this article, we address this question by  
79 presenting three studies comparing Australians' climate-related opinions before and after  
80 the peak of the Black Summer bushfires.

81 **Effects of extreme climate events on climate-related opinions**

82 There are several theoretical and empirical reasons for expecting that extreme  
83 climate events, such as megafires, may prompt shifts in people's climate-related opinions.  
84 First, an often-cited barrier to climate change concern and action is psychological distance

85 (McDonald et al., 2015; Spence et al., 2012; van der Linden et al., 2015; Wang et al., 2019).  
86 It is frequently assumed that many people are unconcerned about climate change because  
87 they are uncertain about whether it is happening (hypothetical distance) and think that, if  
88 it is happening, it will affect other people (social distance), in other places (spatial  
89 distance), in the distant future (temporal distance; for critiques of this idea, see  
90 van Valkengoed et al., 2023; Wang et al., 2018, 2021). Thus, personal experience of  
91 extreme climate events should reduce the psychological distance to climate change,  
92 increasing concern about the issue and willingness to act. Second, although belief in  
93 anthropogenic climate change is generally high amongst the public, there are indications  
94 that the issue is not as salient as other problems (Crawley et al., 2022). Personal  
95 experience of extreme climate events may trigger community discussions that place the  
96 issue “top-of-mind” in the public’s consciousness (Boudet et al., 2020; Demski et al., 2017),  
97 making the problem more salient and increasing support for relevant policies  
98 (Bromley-Trujillo & Poe, 2020). Third, personal experience of extreme climate events  
99 makes the abstract risks of climate change concrete and may provoke negative affective  
100 responses that could increase people’s willingness to mitigate the problem (Bergquist et al.,  
101 2019; Marx et al., 2007; E. U. Weber, 2006). Indeed, it is well-established that the  
102 experience of negative affect associated with climate change is a key predictor of climate  
103 risk perceptions and policy support (Leiserowitz, 2006; van der Linden, 2014, 2015). We  
104 refer to these three mechanisms by which extreme event experiences might influence  
105 climate-related opinions as the psychological distance, issue salience, and affect activation  
106 accounts, respectively (Sisco, 2021).

107 Over the past decade or so, a burgeoning literature has sought to establish whether  
108 personal experience of extreme climate events influences climate-related opinions (for  
109 reviews, see Howe, 2021; Howe et al., 2019; Reser & Bradley, 2020; Reser et al., 2014;  
110 Sambrook et al., 2021; Sisco, 2021). The results of this literature have been somewhat  
111 mixed. On the one hand, and consistent with the precedents just reviewed, several studies

112 have shown that self-reported or objectively recorded personal experience of extreme  
113 climate events, including drought (Carmichael & Brulle, 2017), flooding (Demski et al.,  
114 2017; Ogunbode et al., 2020; Osberghaus & Demski, 2019; Spence et al., 2011; Taylor et al.,  
115 2014), heatwaves (Dai et al., 2015), storms (Bergquist et al., 2019; Lang & Ryder, 2016),  
116 and, notably, forest fires (Lacroix et al., 2020; Zanocco et al., 2018), increases belief in and  
117 concern about climate change. For example, Spence et al. (2011) found that UK households  
118 who reported experiencing flooding were more concerned about climate change, perceived  
119 it as less uncertain, and felt more confident that their actions would make a difference,  
120 compared to households that did not experience flooding. On the other hand, other studies  
121 have failed to document an association between climate-related opinion measures and  
122 personal experience of climate extremes (Cutler et al., 2020; Shao & Hao, 2020), drought  
123 (Carlton et al., 2016), flooding (Albright & Crow, 2019; Whitmarsh, 2008), storms (Lyons  
124 et al., 2018), and multiple disasters including a bushfire, cyclone, and drought (Boon,  
125 2016). A recent meta-analysis using data from 302 studies found that personal experience  
126 of climate events only has a weak positive association with climate change awareness, with  
127 effect sizes varying considerably across different climate events (Xia et al., 2022).

128 There are a few known moderators of the effect of personal experience of extreme  
129 climate events on climate-related opinions (see e.g., Sambrook et al., 2021; Sisco, 2021)  
130 that are especially relevant to the current work. First, personal experience of an extreme  
131 event may only influence climate-related opinions when individuals causally attribute that  
132 event to climate change (E. U. Weber, 2010). Empirical support for this proposition has  
133 been provided in numerous studies (McCright et al., 2014; Ogunbode et al., 2019, 2020;  
134 Wong-Parodi & Rubin, 2022). For example, Ogunbode et al. (2019) found that personal  
135 experience of flooding only predicted climate risk perceptions for individuals who  
136 attributed the flooding to climate change.

137 Second, the effect of personal experience of extreme climate events on  
138 climate-related opinions depends on people's pre-existing climate change beliefs, a pattern

139 consistent with motivated reasoning—the tendency for individuals to interpret information  
140 in ways that protect or reinforce their prior beliefs (Druckman & McGrath, 2019; Kunda,  
141 1990). Exposure to extreme climate events tends to increase climate risk perceptions  
142 primarily among people who already accept climate change, whereas those who are more  
143 sceptical show little or no change (Lacroix et al., 2020). A complementary line of research  
144 uses political ideology as a proxy for these pre-existing climate change beliefs, based on the  
145 well-established association between liberalism and acceptance of climate science and  
146 conservatism and climate scepticism (Hornsey et al., 2016). This work shows that liberals  
147 are more likely than conservatives to perceive or interpret extreme events in ways  
148 consistent with climate change (Lyons et al., 2018; Zanocco et al., 2018). Other work  
149 demonstrates that whether people attribute extreme events to climate change is itself  
150 shaped by their prior beliefs and political ideology (Ogunbode et al., 2019, 2020).

151 Third, media attention to an extreme event may also be necessary for it to shape  
152 climate-related opinions. Extreme climate events can serve as “focusing events” (Birkland,  
153 1998; Birkland & Schwaeble, 2019)—sudden, uncommon, and attention-grabbing  
154 occurrences that attract increased media coverage (Kirilenko et al., 2015; Marquart-Pyatt  
155 et al., 2014; Sisco et al., 2017) and create opportunities to highlight the links between such  
156 events and ongoing climate change. Media attention to climate change has been shown to  
157 influence climate attitudes (Carmichael et al., 2017) and increase public conversations  
158 about the issue (King et al., 2017). However, the effect of media attention may depend on  
159 its frequency and prominence, whether the event is causally attributed to climate change,  
160 and on the presence of competing narratives or misinformation that undermine the  
161 connection between climate change and the extreme event.

## 162 **Divergent mass media and social media bushfire narratives**

163 Because media coverage can shape causal attributions and amplify motivated  
164 reasoning, understanding how the bushfires were portrayed in the media is necessary for  
165 interpreting how Australians made sense of the Black Summer. Mocatta and Hawley

166 (2020) charted the content and evolution of media coverage of the Black Summer, which  
167 focused predominantly on the causes of the fires and what or who was to blame. Scientists  
168 had been quick to confirm that the scale and severity of the fires were unprecedented  
169 (Shine, 2020) and had been worsened by climate change (Climate Council, 2019; Gourlay  
170 et al., 2020). Accordingly, much mass media coverage initially attributed the cause of the  
171 fires to climate change and presented apocalyptic images and descriptions of the  
172 devastation caused. However, as the fires intensified, mass-media reporting of their causes  
173 quickly diverged along ideological lines. Public broadcasters and liberal media outlets  
174 continued to emphasise the connection between climate change and the bushfires, whereas  
175 conservative media outlets sought to downplay the severity of the fires and cast doubt on  
176 the link with climate change. A key argumentation strategy in the conservative media at  
177 this time was the claim that the fires were “nothing new” and in keeping with historic  
178 bushfires in terms of their severity (Johnstone, 2019). Additionally, some conservative  
179 media argued that the Black Summer was worsened by “Greens policies” that prevented  
180 firefighters from reducing fuel loads (G. Brown & Caisley, 2019), despite the Greens’  
181 platform’s overt support for hazard reduction (Australian Greens, 2020).

182       Coinciding with the emergence of these narratives in the conservative mass media,  
183 misinformation began to ferment on the social media platform formally known as Twitter  
184 (now X) in Australia and internationally. Under the hashtag #ArsonEmergency, false  
185 claims began to circulate that the bushfires were caused by arson, that preventative  
186 backburning efforts had been reduced due to green activism, that Australia commonly  
187 experienced such bushfires, and that climate change is unrelated to the bushfires (D. Weber  
188 et al., 2020, 2022). Social media researchers agree that the activities were likely a  
189 deliberate disinformation campaign (Keller et al., 2020; D. Weber et al., 2020).

190       Online misinformation spread under the #ArsonEmergency hashtag, notably the  
191 claim that arsonists were a major cause of the fires, subsequently infected conservative  
192 mass-media reporting of the bushfires. A prominent example was an article published in

193 The Australian under the title “Bushfires: firebugs fuelling crisis as national arson toll hits  
194 183” claiming that “more than 180 alleged arson cases have been recorded since the start of  
195 the bushfire season” (Ross & Reid, 2020). The article played a prominent role in fuelling  
196 online climate change denial narratives and was shared by prominent conservatives, such as  
197 Donald Trump Jr., to his audience of four million followers on Twitter, thus propelling the  
198 misinformation to a much larger online audience. The arson claims were grossly  
199 exaggerated (NSW Bushfire Inquiry, 2020), calculated based on a range of fire-related  
200 offences other than arson, and relied on annual figures rather than the Black Summer  
201 bushfire season (Council, 2021).

202 In summary, media coverage of the Black Summer bushfires focused predominantly  
203 on the causes of the fires and was characterised by a power struggle between two competing  
204 narratives. One narrative emphasised a relationship between climate change and bushfires,  
205 supported by scientists’ assessments of the bushfires (Boer et al., 2020; van Oldenborgh  
206 et al., 2021). The other narrative, fuelled by misinformation, refuted the connection  
207 between climate change and the bushfires, notably by making exaggerated claims about  
208 arson. This polarised and divisive mass media and social media landscape could have  
209 persuaded those undecided about climate change to become more accepting or sceptical  
210 about the issue. Thus, whether the Black Summer bushfires and accompanying media  
211 narratives altered the climate-related opinions of those undecided about climate change is  
212 an open empirical question.

### 213 **Current research**

214 In the following, we report the results of three audience segmentation studies of  
215 Australian climate-related opinions. The studies were undertaken to identify distinct  
216 sub-groups of the Australian population that hold different views about climate change.  
217 All three studies employed the Q methodology (S. R. Brown, 1982; Stephenson, 1986), in  
218 which participants completed a Q-sort task by rank-ordering a series of opinion statements  
219 about climate change, derived from a large-scale analysis of Australian Twitter climate

220 commentary (Andreotta et al., 2019, 2022), according to how similar each was to their  
221 point of view. Participants' rank-orderings of the statements were then subjected to  
222 Q-factor analysis to identify distinct audience segments of climate-related opinions.

223 In Study 1 (September 2019), which took place before the peak of the Black  
224 Summer bushfires, participants completed the Q-sort task along with a battery of measures  
225 of prominent psychological characteristics to help facilitate interpretation of the different  
226 audience segments. We found evidence for a three-segment solution comprising Acceptors,  
227 Fencesitters, and Sceptics, ordered from the highest to the lowest belief in anthropogenic  
228 climate change, trust in climate science, concern about the issue, and motivation to tackle  
229 it. Segments also differed in their climate change concern and scepticism, mental models of  
230 climate change, political ideology, and worldviews, as assessed using the auxiliary  
231 psychological characteristic measures.

232 In Study 2 (February 2020), which took place after the peak of the bushfires,  
233 participants completed the Q-sort task followed by a series of belief-updating tasks to  
234 determine whether segments differed in their receptivity to climate science information. We  
235 replicated the three-segment solution of Study 1 and found considerable heterogeneity in  
236 the belief-updating tendencies of the three segments. Acceptors updated their beliefs the  
237 most towards the scientific estimates, closely followed by Fencesitters, whereas Sceptics did  
238 not update their beliefs at all.

239 These two studies were part of a planned program of research that predated the  
240 bushfires but coincided with their occurrence, affording us a natural experiment, so to  
241 speak, to determine whether the bushfires catalysed a change in Australian climate-related  
242 opinions. The results of these two studies have been reported elsewhere (Andreotta et al.,  
243 2022), but have not yet been systematically compared to determine whether the occurrence  
244 of the bushfires influenced Australian climate-related opinions. In the current paper, we  
245 undertake this comparison and report the results of a third study conducted one month  
246 after our second study, near the end of the Black Summer. In Study 3 (March 2020),

247 participants completed the Q-sort task and the same battery of psychological characteristic  
248 measures used in Study 1. Additionally, participants completed a measure of bushfire  
249 perceptions assessing their endorsement of various media and political claims about the  
250 bushfires—for example, that climate change worsened them, that the fires were severe, and  
251 that arsonists contributed to their occurrence—and a measure of the degree to which the  
252 bushfires warranted a change in Australia’s climate policy.

253 Whereas our earlier research focused on identifying the number, nature, and  
254 psychological characteristics of Australia’s climate change audience segments (Andreotta  
255 et al., 2022), the present work leverages a longitudinal cross-sectional design to examine  
256 whether climate-related opinions shifted across different stages of the Black Summer  
257 bushfires. Using data from all three studies, we first confirmed that the three-segment  
258 solution and the pattern of psychological characteristic differences between segments  
259 reported by Andreotta et al. (2022) generalised to Study 3. Next, we explored whether  
260 climate-related opinions varied in response to the Black Summer bushfires, by testing for  
261 between-study differences in the proportion of respondents assigned to each segment  
262 (Studies 1, 2, & 3) and in climate change cognition and affect (Study 1 vs. Study 3).  
263 Finally, to better understand any observed shifts or stability in climate-related opinions, we  
264 analysed segment-specific perceptions and preferred policy responses to the Black Summer  
265 bushfires (Study 3). Of particular interest was the degree to which respondents causally  
266 attributed the bushfires to climate change versus alternative explanations based on the  
267 erroneous causal claim that the fires were deliberately ignited by arsonists, and whether  
268 these patterns of causal attribution could be understood in terms of motivated reasoning.

269 From an empirical standpoint, the mixed evidence on how extreme climate events  
270 influence climate-related opinions makes it difficult to specify directional hypotheses.  
271 However, from a theoretical standpoint, the psychological distance, issue salience, and  
272 affect activation accounts introduced at the outset predict that—to the extent that people  
273 interpreted the bushfires as having been worsened by climate change—personal experience

274 of the bushfires should make climate change feel more psychologically proximal, heighten  
275 the salience of the issue, and elicit affective responses (e.g., worry) that increase perceived  
276 climate risk. Accordingly, a strong prediction is that the proportion of Fencesitters would  
277 decrease and the proportion of Acceptors would increase across studies. A weaker  
278 prediction is that, even if segment composition remains stable, scores on the continuous  
279 measures of climate change cognition and affect would change in a manner that reflects  
280 greater acceptance of, and concern about, the issue (e.g., increased worry and reduced  
281 scepticism).

282 These effects, however, depend on Australians' perceptions of the causes of the  
283 bushfires. If misinformation narratives gained traction, then individuals may have  
284 attributed the fires to arson rather than to climate change, in which case the psychological  
285 mechanisms identified above would not have been activated. Instead, such narratives could  
286 offset or even reverse the predicted pattern, leading not only to an increase in the  
287 proportion of Fencesitters at the expense of Acceptors, but also to declining concern and  
288 greater scepticism on the continuous climate change cognition and affect measures. On this  
289 misinformation account, climate-related opinions might remain unchanged overall or  
290 instead shift towards greater scepticism.

## 291 Method

292 Data and analysis scripts for this research are available online at  
293 <https://github.com/matt-lab/bushfire-audience-segmentation>. This research was approved  
294 by the Human Research Ethics Committees of the University of Western Australia  
295 (reference: 2019/RA/4/20/5104) and the Commonwealth Scientific and Industrial  
296 Research Organisation (reference: 026/19).

## 297 Participants

298 Table 3 provides an overview of the key characteristics of the study samples and the  
299 materials they completed. Data were collected at three time periods. Study 1 was  
300 conducted in September ( $n = 387$ , 88.97% of Study 1 participants), October ( $n = 42$ ,

301 9.66% of Study 1 participants), and November ( $n = 6$ , 1.38% of Study 1 participants) of  
302 2019, prior to the peak of the Black Summer bushfires. Study 2 was conducted in February  
303 ( $n = 403$ , 97.58% of Study 2 participants) and March ( $n = 10$ , 2.42% of Study 2  
304 participants) of 2020, after the peak of the bushfires. Study 3 was conducted in March  
305 2020 ( $n = 213$ ), approaching the end of the Black Summer bushfires.

306 In total, 1,061 Australian adults participated in the studies. Participants were  
307 recruited using Qualtrics' (Provo, UT) online research panel service using a targeted and  
308 stratified sampling approach to match the age and gender distribution of the general  
309 population (as per the national 2016 census). The age and gender distributions were  
310 comparable across samples, as indicated by regression models (see Supplementary  
311 Methods). These models do not identify statistically significant differences between the  
312 mean age of participants across studies or the (log odds) ratio of female to male  
313 participants across studies. We excluded extremely fast responders who were identified  
314 using a preregistered threshold (see Supplementary Methods).

315 **Materials**

316 ***Q-sort task***

317 To segment participants into climate change audiences, we used the Q-sort task  
318 (S. R. Brown, 1982; Stephenson, 1986). This task requires a set of opinion statements  
319 capturing the breadth of conversational possibilities around an issue, elicited through a  
320 bottom-up process. To create our statements, we drew upon previous work that used an  
321 inductive process to identify the structure of climate change commentary of Australian  
322 tweets (Andreotta et al., 2019). This research revealed five enduring themes of public  
323 discourse on climate change: climate change action, climate change consequences, climate  
324 change conversations, climate change denial, and the legitimacy of climate science and  
325 climate change. For each theme, we selected six tweets that captured the heterogeneity of  
326 the theme (see Andreotta et al., 2022). The resulting 30 tweets were transcribed as  
327 statements that could be understood without the social context of the original tweet.

328 Where possible, language, sentiment, and tone were preserved. Statements included: “It is  
329 important to vote for leaders who will combat climate change” (climate change action),  
330 “Climate change is a threat to the health and safety of our children” (climate change  
331 consequences), “It is shameful that climate change, the greatest problem of our time, is  
332 barely discussed in the media” (climate change conversations), “Climate change sceptics  
333 ignore basic climate science facts” (climate change denial), and “Scientists should stop  
334 falsely claiming that climate change is a settled science” (legitimacy of climate science and  
335 climate change).

336 The Q-sort task was divided into three parts. In part 1, to encourage reflection,  
337 participants began by reading each statement and determining if it was: (1) like their point  
338 of view; (2) unlike their point of view; or (3) neutral or unsure. In part 2, participants  
339 ranked each statement according to how closely it matched their point of view, assigning a  
340 rank from -4 (most unlike their point of view) to +4 (most like their point of view). The  
341 distribution of possible ranks was forced and non-uniform, such that participants had to  
342 consider the few statements to place at the extremes (see Figure 1). This encourages  
343 participants to carefully reflect on their views while completing the task (S. R. Brown,  
344 1982; Stephenson, 1986). Finally, in part 3, participants were asked to justify their  
345 placement of statements assigned extreme ranks.

346 ***Auxiliary psychological scales***

347 A battery of 28 auxiliary psychological characteristic measures was assembled (Table  
348 1). Among these, the most relevant to the current research were state-based psychological  
349 scales of climate change cognition and affect. Specifically, we measured general belief in  
350 anthropogenic climate change, with scales concerning epistemic scepticism (doubt about  
351 anthropogenic climate change), response scepticism (doubt about the effectiveness of  
352 climate change mitigation), perceived human contribution (belief that humans have altered  
353 the global climate), knowledge volume (self-perceived confidence in climate change  
354 knowledge), and worry about climate change. Additionally, we included higher-resolution

355 inventories to quantify participants' mental models of specific climate change causes,  
356 climate change consequences, and the effectiveness of climate change mitigation policies.

357 Other psychological scales pertained to trait-based concepts found to be associated  
358 with climate change beliefs. This includes inventories of: cognitive styles; ideology,  
359 worldviews, and values; and personality.

360 ***Fire Perception Scale***

361 To measure perceptions of the Black Summer bushfires, we developed the Fire  
362 Perception Scale. Drawing on our collective observations of the media and political  
363 discourses surrounding the fires, we identified three prominent themes: climate change  
364 attribution (the extent to which the bushfires were perceived as attributable to climate  
365 change), perceived severity (the perceived magnitude or seriousness of the bushfires), and  
366 human responsibility and preventability (beliefs about whether societal or policy changes  
367 could reduce future fire risk, or whether other human factors such as arson were to blame).

368 We created seven declarative statements reflecting these themes. Example items include  
369 “Climate change made the 2019-20 Australian bushfires more severe” (climate change  
370 attribution), “The 2019-20 Australian bushfires are severe” (perceived severity), and “Over  
371 one hundred arsonists have contributed to the 2019-20 Australian bushfires” (human  
372 responsibility and preventability). Participants rated their agreement with each statement  
373 on a five-point Likert scale: (1) disagree, (2) slightly disagree, (3) neither agree nor  
374 disagree, (4) slightly agree, and (5) agree.

375 ***Policy direction preferences***

376 To measure participants' views on the policy consequences of the Black Summer  
377 bushfires, they responded to two items. First, participants were asked: “Do the 2019-20  
378 Australian bushfires justify a change in Australia's climate change policy?”. Participants  
379 could respond with one of four options: (1) “Yes, the Australian government should be  
380 taking further action to mitigate climate change”; (2) “No, the Australian government  
381 should not modify the current climate change policy”; (3) “Yes, the Australian government

382 should be taking less action to mitigate climate change”; and (4) “Yes, the Australian  
383 government should be taking no action at all to mitigate climate change”. Next,  
384 participants were asked to justify their response (“Why?”) through writing an open-ended  
385 response.

### 386 **Procedure**

387 All studies were conducted as online surveys using Qualtrics (Provo, UT). To begin,  
388 participants read an information sheet, provided informed consent, and supplied basic  
389 demographic information. The procedure subsequently varied across studies (summarised  
390 in Table 3). In Study 1, participants completed the Q-sort task followed by the auxiliary  
391 psychological scales. In Study 2, participants completed the Q-sort task followed by a  
392 belief-updating task unrelated to the current research. In Study 3, participants completed  
393 all materials: the Q-sort task, auxiliary psychological scales, the Fire Perception Scale, and  
394 policy direction preference items. To control for potential order effects, the presentation  
395 sequence of materials was counterbalanced across participants (see Supplementary  
396 Methods).

### 397 **Sample size justification**

398 Sample sizes and the statistical power of our analyses were determined by practical  
399 constraints (Lakens, 2022). Studies 1 ( $n = 435$ ) and 2 ( $n = 413$ ) were undertaken prior to  
400 the current research, with their sample sizes being chosen based on their original objectives  
401 (Andreotta et al., 2022). Study 3 ( $n = 213$ ) was made possible by remaining grant funds  
402 from the earlier studies. Although smaller than the first two studies, this sample  
403 represented the maximum feasible sample size given the available budget and the need for  
404 rapid data collection following the bushfires. To determine the power of tests to detect  
405 study differences in climate change audience segments, cognition, and affect, we conducted  
406 a sensitivity power analysis with the G\*Power program (Faul et al., 2007, 2009). We found  
407 our analyses had sufficient power ( $\geq .80$ ) to detect the expected small effects of study  
408 differences in audience segment membership (for effect sizes of Cohen’s  $\omega \geq 0.106$  for a

409 likelihood-ratio  $\chi^2$  test) and climate change cognition and affect measures (for effect sizes  
410 of Cohen's  $d \geq 0.235$  for  $t$  tests of mean differences).

## 411 Results

412 The results are structured into four sections. First, we assess whether the  
413 three-segment solution and the pattern of psychological characteristic differences between  
414 segments documented in our original analysis of Studies 1 and 2 (Andreotta et al., 2022)  
415 generalise to Study 3. Second, we examine whether the proportion of respondents in each  
416 segment (Studies 1, 2, & 3) and their responses on the climate change cognition and affect  
417 measures (Studies 1 & 3) changed over time. Third, we investigate segment differences in  
418 bushfire perceptions (Study 3), and fourth, we analyse segment differences in policy  
419 preferences (Study 3). All analyses were completed with the *R* programming language (R  
420 Core Team, 2023).

### 421 Replication of the three-segment solution

422 As per our previous research, we used the Q methodology to identify distinct views  
423 on climate change (S. R. Brown, 1982). The Q methodology transposes traditional  
424 dimension reduction techniques to reduce the dimensions of *people* rather than *items*. For  
425 each study, we used principal components analysis with varimax rotation to group  
426 individuals with similar Q-sort ranks. We extracted a single factor, as the second  
427 component accounted for only a minor amount of variance in each study. The extracted  
428 factor represented a dimension of anthropogenic climate-change acceptance. Based on  
429 factor loadings, we divided individuals into one of three segments: (1) *Acceptors* ( $n = 653$ ,  
430 61.55%), whose positive factor loading was statistically significant from zero ( $p < .05$ ); (2)  
431 *Sceptics* ( $n = 97$ , 9.14%), whose negative factor loading was statistically significant from  
432 zero ( $p < .05$ ); and (3) *Fencesitters* ( $n = 311$ , 29.31%), whose factor loading was not  
433 statistically significant from zero ( $p \geq .05$ ).

434 Although the number of segments was consistent across studies, the nature of the  
435 segments may vary. To explore this possibility, we constructed an average Q sort for

436 Acceptors and Sceptics in each study (S. R. Brown, 1982). The ranks assigned to each  
437 statement were averaged (weighted by participants' factor loading). These averages were  
438 then ranked to align with the Q-sort structure, generating a set of values known as factor  
439 scores. For example, the statement with the lowest average corresponded to a factor score  
440 of -4 and the statement with the highest average corresponded to a factor score of +4 (see  
441 Supplementary Results for all factor scores). We did not build a representative Q sort for  
442 Fencesitters as the sorting behaviour of this segment is more heterogenous than the other  
443 two segments (otherwise Fencesitters would have emerged as a separate factor). In all three  
444 studies, the greatest factor score for Acceptors corresponded to the statement "It is  
445 important to vote for leaders who will combat climate change", whereas the greatest factor  
446 score for Sceptics corresponded to the statement "Scientists should stop falsely claiming  
447 that climate change is a settled science."

448 We found minimal differences in each segment's factor scores across studies.

449 Acceptor factor ranks from the three studies were strongly correlated (all Spearman's  $\rho$   
450 correlations  $> .95$ , all  $p$ 's  $< .001$ ). Likewise, Sceptic factor ranks across studies were  
451 strongly correlated (all Spearman's  $\rho$  correlations  $> .94$ , all  $p$ 's  $< .001$ ). Consistently across  
452 studies, Acceptors and Sceptics held divergent views (all Spearman's  $\rho$  correlations  $< -.81$ ,  
453 all  $p$ 's  $< .001$ ). In sum, the number and nature of segments' climate change views were  
454 consistent across time.

455 We also explored whether segments were distinguished by a consistent pattern of  
456 psychological characteristics by replicating the regression analysis of Andreotta et al.  
457 (2022). This analysis was complicated by multicollinearity, which can lead to unstable  
458 coefficient estimates in traditional regression approaches. Instead, we sought to produce  
459 stable estimates with a ridge regression model. A ridge regression reduces the variance of  
460 estimates caused by multicollinearity by shrinking the coefficients towards zero (a  
461 bias-variance tradeoff; James et al., 2021). With the *glmnet* package (Friedman et al.,  
462 2010), we fitted a multinomial logistic ridge regression model to predict segment

463 membership as a function of psychological characteristics for Study 1 and Study 3. The  
464 degree of shrinkage, controlled by a hyperparameter  $\lambda$ , was chosen by a cross-validation  
465 process ( $k$ -fold) that minimised multinomial deviance. Prior to analysis, we converted  
466 responses to  $z$  scores for each predictor in each study. Confidence intervals were estimated  
467 by repeating the modelling procedure via bootstrapping with 10,000 samples (sampled  
468 with replacement; Efron & Tibshirani, 1994).

469 The ridge regression model demonstrated good fit for both Study 1 (83.22%  
470 accuracy, accounting for 49.07% of null deviance) and Study 3 (88.26% accuracy,  
471 accounting for 66.39% of null deviance). As seen in Table 2, the models' coefficients were  
472 generally consistent (same sign) across studies, indicating a robust association between  
473 psychological characteristics and segment membership. Regarding climate change cognition  
474 and affect, Acceptors and Sceptics were distinguished by opposing patterns of climate  
475 change scepticism and belief in anthropogenic climate change. In contrast, the Fencesitters  
476 of Study 3 were characterised by response scepticism and perceptions that carbon-emitting  
477 activities cause climate change. Turning to cognitive styles, conspiracist ideation was  
478 positively associated with Fencesitter membership, and negatively associated with Acceptor  
479 membership (both studies), whereas Sceptics were characterised by a reduced orientation  
480 towards future consequences (Study 3). Generally, Acceptors and Sceptics were  
481 distinguished by opposing patterns of ideologies, worldviews, and values. Lastly,  
482 personality tended not to be a robust predictor of segment membership, although evidence  
483 from Study 3 indicated that Fencesitters were characterised by greater extraversion and  
484 conscientiousness, whereas Sceptics were characterised by greater introversion.

485 **Change in climate change segment membership, cognition, and affect**

486 To explore whether climate change views changed during the Black Summer  
487 bushfires, we investigated the relative proportions of segments across studies (Figure 2).  
488 Numerically, the proportion of Acceptors fell across time (from 64.60% of the Study 1  
489 sample to 54.46% of the Study 3 sample), whereas the proportion of Fencesitters increased

490 across time (from 27.13% of the Study 1 sample to 37.09% of the Study 3 sample). In  
491 comparison, the proportion of Sceptics was relatively stable across studies (from 8.28% of  
492 the Study 1 sample to 8.45% of the Study 3 sample). To investigate whether the relative  
493 proportion of segments differed across studies, we created a multinomial logistic regression  
494 model to predict segment membership as a function of study (coefficients reported in  
495 Supplementary Results), using the *multinom* function from the *nnet* package (Venables &  
496 Ripley, 2002). A likelihood-ratio test did not indicate a statistically reliable improvement  
497 in model fit when study was included as a predictor, compared to a model with only an  
498 intercept term ( $\chi^2 (4) = 8.85, p = .07$ , Cohen's  $\omega = 0.09$ ). Although the effect approached  
499 significance, segment membership did not reliably differ across study samples.

500 In addition to segment membership, we tested for differences in climate change  
501 cognition and affect between Study 1 (September 2019) and Study 3 (March 2020) using *t*  
502 tests. To guard against Type I errors, we applied a Holm (1979) *p* value adjustment (Table  
503 4). Participants in Study 3 reported a significantly higher endorsement of natural cycle  
504 causes of climate change (e.g., volcanic eruptions, solar fluctuations) than those in Study 1  
505 (Cohen's  $d = 0.25$ ). However, no other climate change cognition and affect characteristics  
506 reliably differed between Study 1 and Study 3. Furthermore, there was no evidence that  
507 participants from Study 1 and Study 3 reliably differed in their dispositional attributes of:  
508 cognitive styles; ideology, worldviews and values; or personality (all  $p > .05$ ; see  
509 Supplementary Results for *t* tests).

## 510 **Bushfire perceptions**

511 To explore perceptions of the Black Summer bushfires, we performed a principal  
512 components analysis with varimax rotation on the Fire Perception Scale (see Table 5). We  
513 extracted three factors, as these accounted for the majority of scale variance (78.31%; see  
514 Supplementary Results for scree plot). The first factor, labelled *Climate Processes*, was  
515 characterised by four items (items 1, 3, 5, 6) that linked climate change to the bushfires  
516 and accounted for 41.22% of scale variance. The second factor, labelled *Fire Appraisal*, was

517 characterised by two items (items 2 and 4), with the two most extreme (maximum and  
518 minimum) mean item scores and accounted for 19.97% of scale variance. The third factor,  
519 labelled *Arson Causes*, was characterised by a single item (item 7) stating that the Black  
520 Summer was caused by hundreds of arsonists and accounted for 17.12% of scale variance.  
521 We created subscales corresponding to each factor by averaging item scores. Items that  
522 negatively loaded onto factors were reverse-coded. The multi-item factors of Climate  
523 Processes and Fire Appraisal had an internal consistency of Cronbach's  $\alpha = .86$  (four  
524 items; mean inter-item  $r = .60$ ) and Cronbach's  $\alpha = .42$  (inter-item  $r = .29$ ), respectively.  
525 Given the low internal consistency of Fire Appraisal, we analysed its two component items  
526 separately. For brevity, we refer to these items using abbreviated labels—"Climate change  
527 made bushfires less likely" (item 2) and "Bushfires were severe" (item 4)—in the following  
528 analysis.

529 To test segment differences in bushfire perceptions, we fitted linear regression  
530 models predicting Climate Processes, the two Fire Appraisal items, and Arson Causes from  
531 segment membership (coefficients reported in Supplementary Results). All models  
532 accounted for a significant amount of bushfire perception variance compared to  
533 intercept-only models, indicating that segment membership was a significant predictor of  
534 Climate Processes ( $F(2, 210) = 47.44, p < .001, R^2 = 0.31, R^2_{adjusted} = 0.30$ ), the two Fire  
535 Appraisal items ("Climate change made bushfires less likely":  $F(2, 210) = 18.30, p < .001,$   
536  $R^2 = 0.15, R^2_{adjusted} = 0.14$ ; "Bushfires were severe":  $F(2, 210) = 19.71, p < .001, R^2 =$   
537  $0.16, R^2_{adjusted} = 0.15$ ), and Arson Causes ( $F(2, 210) = 12.69, p < .001, R^2 = 0.11, R^2_{adjusted}$   
538  $= 0.10$ ).

539 To quantify specific segment differences, we conducted pairwise comparisons of  
540 marginal means using the *marginaleffects* package (Arel-Bundock et al., n.d.), with a Holm  
541 (1979)  $p$  value adjustment for multiple comparisons. As seen in Figure 3, Acceptors  
542 reported higher endorsement of Climate Processes than Fencesitters (difference = 0.53,  $SE$   
543 = 0.14, 95% CI = [0.26, 0.80],  $z = 3.87, p < .001, p_{adjusted} < .001$ ), who in turn endorsed

544 Climate Processes more than Sceptics (difference = 1.76,  $SE = 0.25$ , 95% CI = [1.28, 2.24],  
545  $z = 7.14$ ,  $p < .001$ ,  $p_{adjusted} < .001$ ). For the first Fire Appraisal item (“Climate change  
546 made bushfires less likely”), Fencesitters showed higher endorsement than both Acceptors  
547 (difference = 1.04,  $SE = 0.17$ , 95% CI = [0.70, 1.38],  $z = 6.05$ ,  $p < .001$ ,  $p_{adjusted} < .001$ )  
548 and Sceptics (difference = 0.64,  $SE = 0.31$ , 95% CI = [0.04, 1.25],  $z = 2.08$ ,  $p = .037$ ,  
549  $p_{adjusted} = .074$ ). Acceptors and Sceptics did not significantly differ (difference = -0.40,  $SE$   
550 = 0.30, 95% CI = [-0.99, 0.19],  $z = -1.33$ ,  $p = .182$ ,  $p_{adjusted} = .182$ ). For the second Fire  
551 Appraisal item (“Bushfires were severe”), Acceptors reported greater endorsement than  
552 both Fencesitters (difference = 0.65,  $SE = 0.11$ , 95% CI = [0.44, 0.85],  $z = 6.05$ ,  $p < .001$ ,  
553  $p_{adjusted} < .001$ ) and Sceptics (difference = 0.56,  $SE = 0.19$ , 95% CI = [0.20, 0.93],  $z =$   
554 3.04,  $p = .002$ ,  $p_{adjusted} = .005$ ), whereas Fencesitters and Sceptics did not significantly  
555 differ (difference = -0.08,  $SE = 0.19$ , 95% CI = [-0.46, 0.29],  $z = -0.43$ ,  $p = .664$ ,  $p_{adjusted} =$   
556 .664). Finally, for Arson Causes, the pattern of Climate Processes endorsement was  
557 reversed. Specifically, Sceptics showed higher endorsement than Fencesitters (difference =  
558 0.74,  $SE = 0.30$ , 95% CI = [0.15, 1.32],  $z = 2.47$ ,  $p = .014$ ,  $p_{adjusted} = .014$ ), who in turn  
559 showed higher endorsement than Acceptors (difference = 0.55,  $SE = 0.17$ , 95% CI = [0.23,  
560 0.88],  $z = 3.32$ ,  $p < .001$ ,  $p_{adjusted} = .002$ ).

561 We investigated causal perceptions by examining responses to claims that mass  
562 arson (item 7 of the Fire Perception Scale) and climate change (item 1 of the Fire  
563 Perception Scale) contributed to the Black Summer bushfires. The results reveal a striking  
564 pattern of misinformation endorsement. Despite segment differences, participants seldom  
565 rejected the claim that over one hundred arsonists contributed to the Black Summer  
566 bushfires ( $n = 38$ ; 17.84% responded with ‘disagree’ or ‘strongly disagree’ to item 7). In  
567 fact, a majority of Fencesitters (56.96%) and Sceptics (88.89%) endorsed the mass-arson  
568 explanation, compared with 38.79% of Acceptors. This widespread endorsement of an  
569 unsupported causal claim underscores how misinformation narratives surrounding the fires  
570 gained substantial traction—particularly amongst those who are uncertain or sceptical

571 about the scientific consensus on climate change.

572 In contrast, endorsement of the climate-change causal claim showed the opposite  
573 pattern. A majority of Acceptors ( $n = 101$ ; 87.07%) agreed that climate change worsened  
574 the severity of the bushfires, whereas agreement was much lower among Fencesitters  
575 (41.77%) and entirely absent among Sceptics. Overall, endorsement of the mass-arson  
576 causal claim was negatively associated with endorsement of the climate-change causal  
577 claim ( $r = -.21$ , 95% CI =  $[-.33, -.08]$ ,  $p = .002$ ), consistent with the notion that belief in  
578 misinformation displaced acceptance of scientifically grounded explanations.

579 To test a motivated-reasoning account of perceptions of the bushfires, we examined  
580 the association between political liberalism-conservatism (as indexed by the single-item  
581 measure of political ideology)—a proxy for pre-existing climate change beliefs—and  
582 endorsement of the climate-change and mass-arson causal claims. Political liberalism was  
583 associated with stronger endorsement of the climate-change causal claim ( $r = -.34$ , 95% CI  
584 =  $[-0.45, -0.21]$ ,  $p < .001$ ), whereas political conservatism was associated with stronger  
585 endorsement of the mass-arson causal claim ( $r = .27$ , 95% CI =  $[0.15, 0.41]$ ,  $p < .001$ ).  
586 However, both correlations were relatively weak, indicating that the perception that  
587 climate change exacerbated the fires was not tied solely to pre-existing political beliefs, and  
588 that the mass-arson causal claim meshed with a broad range of respondents—including  
589 some who accept anthropogenic climate change—providing only limited support for a  
590 motivated-reasoning explanation (see Supplementary Results for correlations between the  
591 other auxiliary psychological characteristic measures and the Fire Perception Scale causal  
592 items).

### 593 **Policy direction preferences**

594 Policy direction preferences in response to the Black Summer differed among  
595 participants. Most participants desired more governmental climate change mitigation  
596 policies ( $n = 145$ , 68.08%), or no change to existing policies ( $n = 54$ , 25.35%). Only a  
597 small minority desired less or no governmental climate change mitigation policies ( $n = 14$ ,

598 6.57%). However, policy direction preferences differed markedly across segments. The  
599 majority of Acceptors and Fencesitters preferred more mitigation policies, whereas most  
600 Sceptics preferred no change (Figure 4).

601 To test whether these segment differences were statistically reliable, we estimated a  
602 binomial logistic regression model predicting the odds of preferring more mitigation policies  
603 as a function of segment membership (see Supplementary Results for full model outcomes).  
604 Sceptics were excluded from the analysis because none expressed this preference. A  
605 likelihood-ratio test indicated that segment membership significantly predicted policy  
606 direction preferences ( $\chi^2 (1) = 35.45, p < .001$ , Cohen's  $\omega = 0.43$ ). Acceptors  
607 demonstrated a markedly greater preference for more mitigation policies than Fencesitters,  
608 with 89.66% ( $n = 104$ ) endorsing this option compared with 51.90% of Fencesitters ( $n =$   
609 41), corresponding to an odds ratio of 8.03 (95% CI = [3.92, 17.49],  $p < .001$ ).

610 We explored participants' text-based justifications for their policy direction  
611 preferences using an emotion analysis. Emotional associations for each word were detected  
612 using the NRC Word–Emotion Association Lexicon (Mohammad & Turney, 2013), a  
613 crowdsourced lexicon in which words are manually annotated for their association with  
614 eight emotions: anger, fear, anticipation, trust, surprise, sadness, joy, and disgust. For each  
615 response, we coded the presence of each emotion dichotomously (present/not present) if the  
616 response contained at least one word associated with that emotion.

617 The most common emotion expressed by participants was fear ( $n = 67, 31.46\%$ ),  
618 which appeared in justifications for both more action (e.g., “the recent bushfire is a wake-up  
619 call. how much more *worse* do we want to experience?”, fear words italicised) and for no  
620 change or less action (e.g., “...100 arsonists were charged as a starter and it was the fuel left  
621 on the ground for decades that made the fires so much *worse* and caused the *disaster*”).

622 To test whether emotional content varied across segments, we estimated a binomial  
623 logistic regression model for each emotion with segment membership as the predictor (see  
624 Supplementary Results for full model outcomes). We found no statistically significant

625 differences between segments for most emotions, except for fear. Acceptors were more  
626 likely than Fencesitters to use at least one fear-related word, with 40.52% ( $n = 47$ ) doing so  
627 compared with 17.72% of Fencesitters ( $n = 14$ ), corresponding to an odds ratio of 3.16  
628 (95% CI = [1.59, 6.28],  $p = .001$ ). Sceptics fell between these two segments, with 33.33% ( $n$   
629 = 6) using at least one fear-related word, and did not reliably differ from either Acceptors  
630 or Fencesitters.

## 631 Discussion

632 In this paper, we reported three audience segmentation studies of Australian  
633 climate-related opinions employing the Q methodology that were conducted at different  
634 stages of the Australian Black Summer bushfires. Study 1 was conducted before the peak  
635 of the bushfires (September 2019), whereas Studies 2 and 3 took place after the peak  
636 (February and March 2020, respectively). This afforded us a natural experiment to  
637 determine whether the occurrence of the bushfires catalysed a change in Australian  
638 climate-related opinions. All studies required participants to complete a Q-sort task  
639 ranking opinion statements about climate change by degree of endorsement. Studies 1 and  
640 3 additionally incorporated auxiliary measures of prominent psychological characteristics,  
641 including measures of climate change cognition and affect. Study 3 also incorporated  
642 measures of bushfire perceptions and support for climate policy. We examined whether the  
643 three-segment solution and pattern of psychological characteristic differences between  
644 segments reported previously (Andreotta et al., 2022) replicated across studies, whether  
645 the proportion of respondents in each segment and their climate change cognition and  
646 affect differed before versus after the peak of the Black Summer bushfires, and how  
647 segments differed in their bushfire perceptions and policy preferences.

## 648 Summary of key findings

649 Across all three studies, we find consistent support for a three-segment solution of  
650 Australian climate-related opinions. The three segments are the Acceptors, Fencesitters,  
651 and Sceptics—ordered from the highest to the lowest belief in anthropogenic climate

652 change, trust in climate science, concern about the issue, and motivation to tackle it. The  
653 segments are remarkably robust, with near-perfect correlations between the archetypal  
654 sorting styles of Acceptors of all studies and Sceptics of all studies. This is impressive given  
655 that the archetypal sorting styles depend on the correlations between the rank orderings of  
656 30 statements, which have several thousand unique permutations. It appears that the  
657 Black Summer bushfires did not change the definitional point of view of an Acceptor,  
658 Fencesitter, or Sceptic.

659 Further evidence for the stability of audience segments across studies was derived  
660 from the consistent relationship between segments and psychological characteristics in  
661 Studies 1 and 3. Acceptors were characterised by low epistemic and response scepticism,  
662 high worry about climate change, a high belief that carbon-emitting human activities cause  
663 climate change, a high belief in the societal consequences of climate change, a politically  
664 liberal ideology, and an “environment-as-ductile” worldview, meaning they think the  
665 environment has a limited capacity to recover from damage. Sceptics, by contrast, were  
666 characterised by high epistemic and response scepticism, low worry about climate change, a  
667 low belief in the environmental harms of climate change, high confidence in their  
668 knowledge about climate change, a politically conservative ideology, and an  
669 “environment-as-elastic” worldview, meaning they think the environment can easily recover  
670 from damage. In comparison to these two segments, Fencesitters were more neutral  
671 concerning political ideology and environmental worldviews. However, they scored higher  
672 on a measure of general conspiratorial thinking than both Acceptors and Sceptics.

673 We found little evidence to suggest that the Black Summer bushfires catalysed a  
674 shift in climate-related opinions toward greater acceptance and concern. Instead, across the  
675 three studies, the proportion of Acceptors decreased slightly and the proportion of  
676 Fencesitters increased, while Sceptics remained stable. Although the changes in segment  
677 membership over time did not reach conventional levels of statistical significance, the effect  
678 was only marginally non-significant and may reflect the smaller sample size in Study 3

679 compared to Studies 1 and 2.

680 The auxiliary measures of psychological characteristics incorporated in Studies 1 and  
681 3 included several measures of climate change knowledge (viz., knowledge volume, mental  
682 models of climate change, epistemic and response scepticism) and affect (viz., worry about  
683 climate change), affording us an additional set of indicators to determine if the bushfires  
684 provoked a change in beliefs about, and emotional responses towards, climate change.  
685 However, consistent with the results derived from the Q-sort task, we generally found no  
686 statistically reliable change in responses on these measures between Studies 1 and 3. The  
687 only exception was a small increase in Australians' perceptions of natural cycles (e.g.,  
688 volcanic eruptions, solar fluctuations) as a cause of climate change. Again, this evidence  
689 contradicts the claim that the Black Summer bushfires catalysed greater acceptance and  
690 concern about anthropogenic climate change. It is unclear why the Black Summer bushfires  
691 might have strengthened belief in the role of natural cycles in climate change. One  
692 possibility is that participants recognised the greenhouse gases released by the bushfires  
693 and perceived them—along with weather events more broadly—as part of a natural  
694 fluctuation, leading to greater endorsement of natural cycles as a cause of climate change.

695 We did not find any statistically reliable differences between Studies 1 and 3 in a  
696 range of dispositional measures of cognitive style, ideology, worldviews, values, and  
697 personality. This outcome was not unexpected as these are measures of more enduring  
698 psychological traits that tend to remain stable over time.

699 Perceptions of the bushfires in Study 3 varied across segments. Acceptors were most  
700 likely to endorse the view that climate change worsened the fires, followed by Fencesitters,  
701 with Sceptics showing the lowest endorsement. The reverse pattern was observed for the  
702 misinformation-based arson explanation. Sceptics most strongly endorsed the mass-arson  
703 claim, Fencesitters showed moderate endorsement, and even a large minority of Acceptors  
704 endorsed it. Endorsement of the mass-arson explanation was negatively correlated with  
705 endorsement of the climate-change explanation. All segments agreed with the claim that

706 the bushfires were severe, with Acceptors showing the strongest endorsement. However,  
707 Fencesitters were uniquely likely to endorse the counterfactual claim that climate change  
708 made the bushfires less likely—more so than both Acceptors and Sceptics.

709 Finally, support for climate policy in Study 3 also varied by segments. Acceptors  
710 almost universally agreed that the bushfires warranted more action by Australia to address  
711 climate change, whereas Fencesitters were roughly evenly split between favouring more  
712 action and no change in action. Sceptics mostly favoured no change in action by Australia  
713 to address climate change. Fear was routinely used by all segments, but in particular,  
714 Acceptors, to justify their policy position.

### 715 **Why the Black Summer did not lead to greater climate change concern**

716 Our results add to the mixed findings on the relationship between climate-related  
717 opinions and personal experience of extreme climate events (Howe, 2021; Howe et al., 2019;  
718 Xia et al., 2022). At the outset, we identified three moderators of the effect of  
719 extreme-event exposure on climate-related opinions—event attribution, motivated  
720 reasoning, and media attention (Sisco, 2021). Taken together, these moderators may offer a  
721 coherent explanation of why the Black Summer bushfires did not lead to greater  
722 acceptance and concern about climate change.

723 The first and most important moderator is event attribution. Previous work shows  
724 that extreme-event exposure influences climate-related opinions only among individuals  
725 who believe the event was caused by climate change (McCright et al., 2014; Ogunbode  
726 et al., 2019, 2020; Wong-Parodi & Rubin, 2022). In Study 3, this key precondition was not  
727 met for the vast majority of respondents—and critically, it was not achieved for  
728 Fencesitters, the segment most theoretically open to changing their opinions in response to  
729 personal experience. Most Fencesitters and all Sceptics rejected the notion that climate  
730 change contributed to the fires and instead attributed the bushfires primarily to arson.  
731 Although Acceptors were more likely to endorse a climate-change explanation, even within  
732 this segment, a sizeable minority also attributed the fires to arson. Thus, because most

733 respondents—including the Fencesitters most amenable to opinion change—did not  
734 interpret the fires as climate-related, the chief mechanism that links extreme-event  
735 exposure to increased climate change concern was not engaged.

736 The second moderator, motivated reasoning, partially helps to explain why these  
737 attribution patterns emerged. Consistent with previous research (Ogunbode et al., 2019,  
738 2020), political ideology influenced causal attributions—liberals were more likely to  
739 attribute the fires to climate change, whereas conservatives were more likely to endorse the  
740 arson explanation. However, motivated reasoning alone cannot fully account for the  
741 findings, as belief in the arson narrative was evident across all three segments—including  
742 among some Acceptors—suggesting that wider informational influences were at play.

743 The third moderator that captures these wider informational influences is media  
744 attention. Extreme-climate events can serve as “focusing events” (Birkland, 1998; Birkland  
745 & Schwaebel, 2019), increasing public and media attention and generating opportunities to  
746 highlight the links between such events and climate change (Kirilenko et al., 2015;  
747 Marquart-Pyatt et al., 2014; Sisco et al., 2017). Although the Black Summer bushfires  
748 received substantial coverage, the media environment was highly contested. After an initial  
749 period emphasising the role of climate change, the narrative quickly became politicised  
750 (Mocatta & Hawley, 2020). Liberal media outlets continued to stress the role of climate  
751 change in worsening the fires, whereas conservative outlets were dismissive of this  
752 connection. During the peak of the bushfires, misinformation—particularly exaggerated  
753 claims about widespread arson—proliferated across social media and conservative media  
754 outlets. Such misinformation is notoriously difficult to correct (Ecker et al., 2022;  
755 Lewandowsky et al., 2012), and it is possible that despite officials’ efforts to dismiss the  
756 arson claims (Knaus, 2020; Readfearn, 2019), these narratives had already taken hold by  
757 the time Studies 2 and 3 were conducted. The widespread acceptance of the mass-arson  
758 explanation among Fencesitters and Sceptics, in particular, is consistent with this  
759 interpretation. It may also help explain why the proportion of Acceptors decreased while

760 the proportion of Fencesitters increased over time. Misinformation exposure may have  
761 rendered some Acceptors more sceptical of the scientific consensus on climate change.

762 However, misinformation is not the only media-related explanation. A content  
763 analysis of Australian news coverage by Burgess et al. (2020) found that although nearly  
764 half of articles mentioned climate change, only 16% explicitly attributed the bushfires to  
765 climate change—even though climate scientists and scientific bodies made this connection  
766 during the fires (Climate Council, 2019; Hughes et al., 2020; Readfearn, 2020; World  
767 Weather Attribution Consortium, 2020)—and even fewer explained how climate change  
768 contributed to the fires. Similar results were observed in a study examining how Australian  
769 climate-action non-governmental organisations framed the link between the Black Summer  
770 bushfires and climate change on the social media platform Twitter (now X; Ettinger et al.,  
771 2023). These results indicate that climate-communication stakeholders may not have  
772 provided the public with sufficiently clear explanations of the relationship between climate  
773 change and bushfire risk, limiting opportunities for accurate event attribution.

774 Finally, Studies 2 and 3 were undertaken after the World Health Organisation  
775 (WHO) declared the COVID-19 outbreak a Public Health Emergency of International  
776 Concern in January 2020, and Study 3 coincided with the WHO characterising the  
777 outbreak as a pandemic in March 2020. The abrupt nature of the pandemic meant that it  
778 quickly became the centre of global media and public attention, diverting the spotlight  
779 away from the bushfires and climate change (Evensen et al., 2021; Loureiro & Alló, 2021;  
780 Rauchfleisch et al., 2023; Smirnov & Hsieh, 2022; Stoddart et al., 2023). Accordingly, the  
781 absence of a shift towards greater climate-change concern may reflect a redirection of  
782 attention and worry towards the rapidly escalating global health crisis.

783 It is important to acknowledge that these explanations remain tentative, as our  
784 study was not a true experiment. Specifically, we lack the relevant counterfactual  
785 conditions to facilitate causal inference—for example, a less polarised media environment,  
786 an absence of misinformation, or a scenario in which the bushfires did not coincide with a

787 global health crisis.

788 **Implications for theoretical accounts of the effects of extreme-event experiences**

789 At first blush, the results appear at variance with the psychological distance, issue  
790 salience, and affect activation accounts introduced at the outset. According to these  
791 accounts, personal experience of the bushfires should have made climate change feel more  
792 psychologically proximal, heightened the salience of the issue, and elicited affective  
793 responses that increased perceived climate risk. Consequently, we would expect to observe  
794 an increase in the proportion of Acceptors and a decrease in the proportion of Fencesitters  
795 over time, and/or shifts on the continuous climate change cognition and affect measures  
796 towards greater acceptance and concern about the issue. Yet, contrary to these predictions,  
797 we found a slight decrease in Acceptors and a corresponding increase in Fencesitters, and  
798 no reliable changes on the continuous cognition and affect measures.

799 However, these findings must be interpreted in light of the moderating variables just  
800 considered. A tacit assumption of the above theoretical accounts is that people perceived  
801 the bushfires as a causal consequence of climate change or as an event exacerbated by  
802 climate change. As our results show, this assumption did not hold for a substantial  
803 proportion of respondents. Many attributed the bushfires primarily to mass arson causes,  
804 and endorsement of this claim was associated with rejection of the climate-change causal  
805 explanation. Accordingly, the key preconditions required for the psychological distance,  
806 issue salience, and affect activation accounts to generate the predicted changes in  
807 climate-related opinions were unlikely to have been met. This limits the extent to which  
808 strong inferences can be drawn about the validity of these theories in this context.

809 Beyond these psychological accounts, the findings have implications for the  
810 focusing-events framework (Birkland, 1998; Birkland & Schwaebel, 2019). Extreme events  
811 can serve as focusing events that increase public attention to an issue and create  
812 opportunities for attitude change. However, the Black Summer bushfires did not appear to  
813 have this effect. Media narratives surrounding the fires quickly became contested and

814 politicised, with climate-change interpretations competing against misinformation-based  
815 arson narratives, while the emerging COVID-19 crisis eventually hijacked the spotlight of  
816 public and media attention. This suggests that focusing events are vulnerable to  
817 politicisation, alternative causal narratives, and competing, attention-grabbing events.  
818 Such conditions may prevent the expected increase in climate-change concern or even  
819 produce shifts toward greater scepticism if misinformation about the event dominates  
820 media and public discourse.

### 821 **Implications for climate change communication**

822 Our results have implications for how climate-communication stakeholders frame  
823 extreme climate events. Providing clear statements attributing such events to climate  
824 change is important, given evidence that event attribution is a key moderator of the effect  
825 of extreme-event exposure on climate-related opinions. However, an even more critical step  
826 is to explain, in simple terms, how climate change contributed to the event. Doing so  
827 makes the causal claim more credible and memorable because the underlying mechanism is  
828 understood (Hastie, 1984) and may help to stave off misconceptions fuelled by  
829 misinformation. For instance, if individuals understand that climate change did not ignite  
830 the fires but rather created the unusually hot and dry conditions that enabled them to  
831 burn intensely and spread rapidly, they may be less likely to be swayed by claims that the  
832 fires were caused by arson rather than climate change. In other words, they will recognise  
833 that the source of ignition is largely inconsequential—climate change does not start  
834 bushfires, rather it creates conditions that worsen them once ignited. It is clear from the  
835 content analyses of mass media and social media coverage of the Black Summer by Burgess  
836 et al. (2020) and Ettinger et al. (2023), respectively, described earlier, that such  
837 explanatory links were frequently missing—even though climate scientists and scientific  
838 bodies explicitly connected the fires to climate change at the time (Climate Council, 2019;  
839 Hughes et al., 2020; Readfearn, 2020; World Weather Attribution Consortium,  
840 2020)—suggesting that more could have been done to communicate clearly how climate

841 change exacerbated the fires.

842 When misinformation about the causes of an extreme climate event circulates in  
843 mass and social media, timely correction may be crucial to prevent it from taking root.  
844 During the Black Summer bushfires, state fire services, the police, and journalists all  
845 played roles in countering misinformation about the fires. However, not all corrections are  
846 equally effective in debunking misinformation. Cognitive psychologists have identified  
847 numerous best practices for debunking misinformation (Ecker et al., 2022; Lewandowsky  
848 et al., 2012), and these strategies have been distilled into an accessible handbook for  
849 non-experts (Lewandowsky et al., 2020). A core principle is that effective corrections  
850 should not simply retract a false claim but provide a plausible alternative explanation for  
851 the cause of an event (Ecker et al., 2022). For instance, when countering the claim that  
852 “the bushfires were caused by arsonists,” a correction that offers an alternative causal  
853 account (e.g., “the bushfires were ignited by lightning”) is more effective than one that  
854 merely states that the claim is false (e.g., “there is no evidence of arson”).

855 Climate-communication stakeholders should incorporate these best-practice insights into  
856 their messaging to increase the effectiveness of their debunking efforts. Even members of  
857 the public can help limit the spread of misinformation. For example, in their analysis of  
858 the #ArsonEmergency tweets on Twitter, Weber and colleagues (D. Weber et al., 2020,  
859 2022) identified two different communities, one involved in the propagation of the false  
860 claims and another that sought to debunk those claims.

861 Finally, our results have implications for engaging with the three audience segments.  
862 Acceptors and Sceptics may be relatively low priorities for public engagement campaigns,  
863 albeit for different reasons. Acceptors already strongly believe in anthropogenic climate  
864 change, are highly trusting of climate science, and are strongly supportive of climate  
865 action. Accordingly, additional messaging is unlikely to appreciably shift their views.  
866 Sceptics, by contrast, have firmly entrenched beliefs. Their conservative political ideology  
867 and environment-as-elastic worldviews render them motivated to discount climate science,

meaning they are highly resistant to belief revision in the face of climate science information (Andreotta et al., 2022). Combined with the fact that they are few in number, these characteristics suggest there may be little merit in trying to shift the opinions of this segment (although see Andreotta et al., 2022, for a more nuanced account). By comparison, Fencesitters represent a more promising group for public engagement. They are more neutral in terms of political ideology and environmental worldviews, meaning they are not politically motivated to oppose climate science, in contrast to Sceptics. Indeed, Fencesitters update their beliefs in response to climate science information almost as much as Acceptors do (Andreotta et al., 2022). They are a relatively large segment with more intermediate climate-related opinions, meaning that with the right messaging strategy, they could perhaps be transformed into Acceptors.

Accordingly, we suggest that public engagement campaigns should target the Fencesitters. Unfortunately, we know little about the characteristics of this segment. This is, in part, because, given the inherent variability of individuals within this segment, we cannot construct an archetypal Q sort of their statement rankings. However, what we do know is that, compared to the Acceptors and Sceptics, they are more likely to endorse conspiracy theories<sup>1</sup>. This curious result, first documented in our original report of Studies 1 and 2 (Andreotta et al., 2022), was replicated in Study 3, suggesting it is a robust feature of this segment. Given that much climate misinformation is grounded in conspiracy theories (Coan et al., 2021; Cook, 2020), our key piece of advice for climate-communication stakeholders is that debunking efforts should pay particular attention to exposing how climate misinformants employ conspiracy theories and related deception techniques to mislead the public. Such refutation techniques may be necessary to prevent climate

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<sup>1</sup> It is unclear why Fencesitters scored higher on conspiracist ideation than Acceptors and Sceptics. If anything, we would have expected Sceptics to score higher on this psychological attribute, based on previous evidence showing an association between conspiracy theory endorsement and climate-change scepticism (Lewandowsky et al., 2013). One possibility is that the heightened responses of Fencesitters on this measure reflect a broader psychological attribute, such as openness to alternative viewpoints—the willingness to consider alternative ideas, suggestions, or explanations of events (Tsai & Li, 2023).

891 misinformation from transforming Fencesitters into Sceptics.

892 **Potential limitations**

893 Before closing, some potential limitations of the current work merit comment. First,  
894 Studies 2 and 3 were undertaken after the peak in the bushfires, which occurred between  
895 December 2019 and January 2020. Therefore, we cannot rule out the possibility that had  
896 data been collected during the most intense period of the crisis, a transient increase in  
897 climate change acceptance and concern may have been detected. Nevertheless, even if this  
898 were so, our longitudinal comparison suggests such a change in opinions would have been  
899 temporary and short-lived.

900 Second, although around 80% of the Australian population was affected either  
901 directly or indirectly by the fires (Hughes et al., 2020), we did not ask respondents about  
902 the nature of their experiences. The distinction between direct and indirect experience is  
903 important because studies have shown that direct experience of an extreme event is more  
904 predictive of climate-related opinions than indirect experience (Ogunbode et al., 2020;  
905 Zanocco et al., 2019). Accordingly, changes in climate-related opinions are more likely to be  
906 observed among individuals who had severe direct negative experiences of the fires, such as  
907 those who suffered property damage. However, we note that even if we had measured the  
908 nature of our respondents' experiences, individuals who had a severe personal experience of  
909 the fires are likely under-represented in Studies 2 and 3, as the disaster's impact would  
910 have precluded them from responding to our web-panel surveys (Howe, 2021).

911 Third, and relatedly, we did not collect respondents' state or region of residence.

912 Because the timing, intensity, and immediacy of the bushfires varied across Australia,  
913 regional variation could have shaped respondents' levels of exposure and perceptions.  
914 Although indirect exposure—in the form of smoke, extensive media coverage, and national  
915 political debate—was widespread across the country, the absence of geographic information  
916 means we cannot rule out the possibility that regional differences contributed to some of  
917 the observed trends.

918 Finally, we note that power is always a concern when retaining the null hypothesis.

919 Our sample size had sufficient power to detect even very weak omnibus effects, such as  
920 study differences in segment membership and climate change cognition and affect. We had  
921 less power to detect post hoc effects between specific segments and specific studies, such as  
922 the increase in Fencesitters between Study 1 and Study 3. However, our studies detected  
923 some key segment differences, such as Fencesitters endorsing arson-based explanations of  
924 the Black Summer bushfires at a greater rate than Acceptors.

## 925 **Conclusions**

926 Previous research examining the association between personal experience of extreme

927 climate events and climate-related opinions has produced contradictory findings. It is  
928 therefore perhaps unsurprising that we found no evidence that the Black Summer bushfires  
929 prompted a shift towards greater acceptance of, and concern about, anthropogenic climate

930 change. A key determinant of whether extreme-event exposure influences climate-related  
931 opinions is whether people attribute the event to climate change. Accordingly, the failure  
932 of Fencesitters to attribute the bushfires to climate change is the most credible explanation  
933 for the lack of a positive shift in climate-related opinions. This lack of attribution may

934 reflect the divergent mass-media and social-media narratives surrounding the bushfires. In  
935 particular, misinformation dismissing the connection between climate change and the

936 bushfires may have “crowded out” messages linking the fires to climate change. That such

937 misinformation influenced perceptions is evidenced by the sizeable minority of Acceptors

938 who did not reject the arson claim, and by the majority of Fencesitters and Sceptics who

939 endorsed it. Although some mass-media coverage mentioned climate change, relatively few

940 articles explicitly linked the bushfires to climate change, and fewer still explained the

941 mechanism by which climate change intensifies bushfire conditions—another likely reason

942 why attribution rates were low among Fencesitters. Motivated reasoning also played a role,

943 but its influence appeared limited. While politically conservative respondents were more

944 likely to endorse the arson explanation, belief in the mass-arson narrative was found across

945 all three segments, indicating that misinformation narratives exerted widespread influence.

946 The implications of these observations are twofold. First, climate-communication

947 stakeholders may need to emphasise not only the connection between an extreme event and

948 climate change but also explain clearly how climate change contributed to that event.

949 Second, where misinformation about the cause of an extreme event circulates, proactive

950 efforts must be undertaken to debunk the misleading claims. This requires that

951 climate-communication stakeholders are aware of best practices for refuting misinformation

952 so that their interventions can achieve maximal impact.

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**Table 1***Summary of auxiliary psychological measures.*

Psychological characteristic	Items	Cronbach's α	Range	Example item	Reference
<b>Climate change cognition and affect</b>					
Knowledge Volume	1	-	1 to 4	How much do you feel you know about climate change?	Malka et al. (2009)
Perceptions of Carbon Emission	7	0.92	1 to 7	Please indicate to what extent each of the following is a cause of climate change, to the best of your knowledge: people driving their cars	Andreotta et al. (2022)
Causes					
Perceptions of Environmental Harm	4	0.87	1 to 7	Please indicate to what extent each of the following is a cause of climate change, to the best of your knowledge: air pollution from toxic chemicals	Andreotta et al. (2022)
Causes					
Perceptions of Natural Causes	2	0.79	1 to 7	Please indicate to what extent each of the following is a cause of climate change, to the best of your knowledge: volcanic eruptions	Andreotta et al. (2022)
Perceived Personal Consequences	3	0.87	1 to 7	Please rate for each of the following how likely it is as a consequence of climate change by the year 2050: food shortages where you live	Bostrom et al. (2012)
Perceived Societal Consequences	8	0.96	1 to 7	Please rate for each of the following how likely it is as a consequence of climate change by the year 2050: food shortages in many parts of the world	Bostrom et al. (2012)
Perceived Human Contribution	1	-	1 to 7	How likely do you think it is that human actions have changed global climate?	Bostrom et al. (2012)
Perceived Effectiveness of Carbon Policies	3	0.75	1 to 7	Please rate for each step what effect you think it would have on climate change: requiring cars and trucks to have higher fuel efficiency (1 = Reduce or Stop Climate Change, 4 = Neither Reduce nor Increase, 7 = Increase Climate Change)	Bostrom et al. (2012)

(continued)

Psychological characteristic	Items	Cronbach's α	Range	Example item	Reference
Perceived Effectiveness of Engineering Policies	3	0.42	1 to 7	Please rate for each step what effect you think it would have on climate change: putting more dust in the atmosphere (1 = Reduce or Stop Climate Change, 4 = Neither Reduce nor Increase, 7 = Increase Climate Change)	Bostrom et al. (2012)
Perceived Effectiveness of Green Policies	5	0.91	1 to 7	Please rate for each step what effect you think it would have on climate change: planting trees (1 = Reduce or Stop Climate Change, 4 = Neither Reduce nor Increase, 7 = Increase Climate Change)	Bostrom et al. (2012)
Epistemic Scepticism	8	0.91	1 to 5	Climate change is just a natural fluctuation in Earth's temperatures	Capstick and Pidgeon (2014)
Response Scepticism	7	0.89	1 to 5	There is no point in me doing anything about climate change because no-one else is	Capstick and Pidgeon (2014)
Worry about Climate Change	1	-	1 to 4	How strongly do you feel worry when you think about the issue of climate change?	Smith and Leiserowitz (2014)
<b>Cognitive style</b>					
Orientation to Future Goals	4	0.72	1 to 5	I consider how things might be in the future	Enzler (2015)
Orientation to Immediate Goals	5	0.86	1 to 5	I mainly act to satisfy my immediate concerns, figuring the future will take care of itself	Enzler (2015)
Conspiracist Ideation	6	0.90	1 to 5	The Apollo moon landings never happened and were staged in a Hollywood film studio	Lewandowsky et al. (2013)
Need for Cognition	6	0.79	1 to 5	I would prefer complex to simple problems	Lins de Holanda Coelho et al. (2018)
<b>Ideology, worldviews, and values</b>					
Environment-as-Ductile Worldview	6	0.81	1 to 5	If the balance of the natural environment is upset the whole system will collapse	Price et al. (2014)
Environment-as-Elastic Worldview	6	0.85	1 to 5	The natural environment is capable of recovering from any damage humans may cause	Price et al. (2014)

*(continued)*

Psychological characteristic	Items	Cronbach's $\alpha$	Range	Example item	Reference
Political Ideology	1	-	1 to 7	Please indicate the extent to which you identify yourself as - politically left-wing or right-wing (1 = Very Left-Wing, 7 = Very Right-Wing)	-
System Justification	8	0.85	1 to 9	Everyone has a fair shot at wealth and happiness	Kay and Jost (2003)
Conservation Values	10	0.32	-2.94 to 5.54	Please, rate the importance of the following values as a life-guiding principle for you: CONFORMITY (obedience, honouring parents and elders, self-discipline, politeness)	Lindeman and Verkasalo (2005)
Self-Transcendence Values	10	0.55	-4.84 to 2.52	Please, rate the importance of the following values as a life-guiding principle for you: BENEVOLENCE (helpfulness, honesty, forgiveness, loyalty, responsibility)	Lindeman and Verkasalo (2005)
<b>Personality</b>					
Agreeableness	2	0.27	1 to 5	I see myself as someone who is generally trusting	Rammstedt and John (2007)
Conscientiousness	2	0.53	1 to 5	I see myself as someone who does a thorough job	Rammstedt and John (2007)
Extraversion	2	0.53	1 to 5	I see myself as someone who is outgoing, sociable	Rammstedt and John (2007)
Neuroticism	2	0.62	1 to 5	I see myself as someone who gets nervous easily	Rammstedt and John (2007)
Openness	2	0.14	1 to 5	I see myself as someone who has an active imagination	Rammstedt and John (2007)

*Note:*

Conservation and Self-Transcendence Value scores were a weighted average of ten items (rated along a nine-point scale). Table reproduced with updated Cronbach's  $\alpha$  from Andreotta et al. (2022), under the Creative Commons license (CC BY 4.0).

**Table 2**

*Effect of psychological characteristics on segment membership, as estimated by a multinomial logistic ridge regression for Studies 1 and 3.*

Predictors	Acceptors		Fencesitters		Sceptics	
	Study 1	Study 3	Study 1	Study 3	Study 1	Study 3
<b>Intercept</b>						
	+1.64 <sup>^</sup> [1.64, 2.18]	+1.66 <sup>^</sup> [1.44, 2.09]	+0.56 <sup>^</sup> [0.44, 0.99]	+1.03 <sup>^</sup> [0.71, 1.32]	-2.20 <sup>^</sup> [-3.06, -2.19]	-2.69 <sup>^</sup> [-3.22, -2.36]
<b>Climate change cognition and affect</b>						
Epistemic Scepticism	-0.33 <sup>^</sup> [-0.59, -0.26]	-0.46 <sup>^</sup> [-0.72, -0.25]	+0.11 [-0.05, 0.30]	+0.13 [-0.08, 0.39]	+0.23 <sup>^</sup> [0.16, 0.43]	+0.33 <sup>^</sup> [0.19, 0.46]
Worry about Climate Change	+0.31 <sup>^</sup> [0.23, 0.60]	+0.13 [-0.09, 0.38]	-0.06 [-0.25, 0.11]	+0.10 [-0.12, 0.36]	-0.25 <sup>^</sup> [-0.50, -0.19]	-0.23 <sup>^</sup> [-0.44, -0.07]
Response Scepticism	-0.29 <sup>^</sup> [-0.55, -0.19]	-0.55 <sup>^</sup> [-0.75, -0.37]	+0.08 [-0.09, 0.28]	+0.34 <sup>^</sup> [0.14, 0.56]	+0.21 <sup>^</sup> [0.15, 0.40]	+0.21 <sup>^</sup> [0.09, 0.35]
Perceived Human Contribution	+0.20 <sup>^</sup> [0.08, 0.41]	+0.27 <sup>^</sup> [0.12, 0.51]	+0.12 [-0.02, 0.35]	-0.06 [-0.29, 0.16]	-0.32 <sup>^</sup> [-0.59, -0.23]	-0.22 <sup>^</sup> [-0.42, -0.07]
Perceived Societal Consequences	+0.19 <sup>^</sup> [0.06, 0.39]	+0.11 [-0.08, 0.38]	-0.09 [-0.30, 0.05]	+0.06 [-0.21, 0.25]	-0.10 [-0.23, 0.04]	-0.16 <sup>^</sup> [-0.33, -0.02]
Perceptions of Environmental Harm Causes	+0.08 [-0.09, 0.26]	+0.04 [-0.18, 0.24]	+0.08 [-0.08, 0.28]	+0.19 [0.00, 0.43]	-0.16 <sup>^</sup> [-0.32, -0.05]	-0.22 <sup>^</sup> [-0.37, -0.10]
Knowledge Volume	-0.10 [-0.34, 0.01]	-0.05 [-0.25, 0.13]	-0.06 [-0.24, 0.10]	-0.01 [-0.22, 0.19]	+0.15 <sup>^</sup> [0.04, 0.43]	+0.06 [-0.11, 0.26]
Perceptions of Carbon Emission Causes	+0.15 <sup>^</sup> [0.00, 0.35]	+0.15 [-0.02, 0.32]	+0.04 [-0.11, 0.23]	+0.29 <sup>^</sup> [0.09, 0.49]	-0.19 <sup>^</sup> [-0.36, -0.11]	-0.44 <sup>^</sup> [-0.59, -0.29]
Perceived Effectiveness of Engineering Policies	-0.13 <sup>^</sup> [-0.36, -0.01]	+0.09 [-0.11, 0.31]	+0.14 <sup>^</sup> [0.01, 0.36]	-0.10 [-0.31, 0.11]	-0.01 [-0.14, 0.15]	+0.01 [-0.15, 0.16]
Perceived Personal Consequences	+0.12 [-0.03, 0.30]	+0.12 [-0.09, 0.36]	-0.02 [-0.19, 0.14]	-0.09 [-0.31, 0.15]	-0.10 [-0.23, 0.02]	-0.03 [-0.21, 0.11]
Perceived Effectiveness of Carbon Policies	+0.11 <sup>^</sup> [-0.03, 0.35]	-0.13 [-0.34, 0.09]	-0.03 [-0.23, 0.15]	+0.17 [-0.07, 0.36]	-0.08 [-0.27, 0.02]	-0.03 [-0.16, 0.13]
Perceived Effectiveness of Green Policies	+0.10 [-0.02, 0.30]	-0.04 [-0.24, 0.17]	-0.04 [-0.20, 0.14]	+0.10 [-0.12, 0.31]	-0.06 [-0.27, 0.05]	-0.06 [-0.20, 0.08]
Perceptions of Natural Causes	-0.08 [-0.26, 0.08]	-0.15 [-0.40, 0.05]	+0.05 [-0.10, 0.24]	+0.10 [-0.12, 0.36]	+0.02 [-0.15, 0.20]	+0.05 [-0.16, 0.25]
<b>Cognitive style</b>						
Orientation to Future Goals	+0.05 [-0.11, 0.25]	+0.21 [0.00, 0.38]	+0.06 [-0.10, 0.26]	+0.10 [-0.09, 0.30]	-0.11 [-0.33, 0.04]	-0.31 <sup>^</sup> [-0.47, -0.11]
Conspiracist Ideation	-0.15 <sup>^</sup> [-0.36, -0.02]	-0.49 <sup>^</sup> [-0.70, -0.32]	+0.15 <sup>^</sup> [0.02, 0.36]	+0.33 <sup>^</sup> [0.15, 0.55]	+0.00 [-0.18, 0.17]	+0.16 [-0.02, 0.34]
Need for Cognition	-0.12 [-0.32, 0.01]	-0.07 [-0.25, 0.15]	+0.01 [-0.15, 0.18]	-0.02 [-0.23, 0.19]	+0.10 [-0.03, 0.31]	+0.09 [-0.12, 0.27]
Orientation to Immediate Goals	+0.02 [-0.12, 0.25]	-0.16 [-0.42, 0.00]	0.00 [-0.20, 0.17]	+0.15 [-0.04, 0.41]	-0.02 [-0.21, 0.10]	+0.02 [-0.18, 0.21]

(continued)

Predictors	Acceptors		Fencesitters		Sceptics	
	Study 1	Study 3	Study 1	Study 3	Study 1	Study 3
<b>Ideology, worldviews, and values</b>						
Environment-as-Ductile Worldview	+0.18 [-0.01, 0.44]	<b>+0.40<sup>^</sup></b> [ <b>0.23, 0.62</b> ]	-0.11 [-0.36, 0.05]	<b>-0.21<sup>^</sup></b> [-0.43, -0.01]	-0.07 [-0.21, 0.10]	<b>-0.19<sup>^</sup></b> [-0.36, -0.04]
Conservation Values	-0.11 [-0.32, 0.02]	<b>-0.26<sup>^</sup></b> [-0.46, -0.06]	+0.01 [-0.17, 0.18]	-0.02 [-0.22, 0.22]	+0.11 [-0.05, 0.32]	<b>+0.27<sup>^</sup></b> [0.06, 0.45]
Environment-as-Elastic Worldview	<b>-0.20<sup>^</sup></b> [-0.43, -0.05]	<b>-0.37<sup>^</sup></b> [-0.58, -0.20]	+0.05 [-0.15, 0.23]	+0.07 [-0.12, 0.33]	<b>+0.15<sup>^</sup></b> [0.03, 0.38]	<b>+0.30<sup>^</sup></b> [0.12, 0.46]
System Justification	+0.04 [-0.12, 0.25]	<b>+0.20<sup>^</sup></b> [0.04, 0.39]	+0.06 [-0.12, 0.23]	<b>-0.23<sup>^</sup></b> [-0.44, -0.04]	-0.09 [-0.30, 0.07]	+0.03 [-0.16, 0.22]
Self-Transcendence Values	+0.04 [-0.10, 0.21]	+0.17 [-0.04, 0.36]	-0.10 [-0.28, 0.05]	+0.02 [-0.20, 0.21]	+0.06 [-0.12, 0.24]	<b>-0.19<sup>^</sup></b> [-0.33, 0.00]
Political Ideology	<b>-0.18<sup>^</sup></b> [-0.41, -0.04]	-0.10 [-0.35, 0.12]	+0.03 [-0.17, 0.19]	-0.16 [-0.38, 0.06]	<b>+0.16<sup>^</sup></b> [0.02, 0.40]	<b>+0.26<sup>^</sup></b> [0.09, 0.47]
<b>Personality</b>						
Extraversion	-0.01 [-0.15, 0.14]	+0.03 [-0.21, 0.22]	+0.03 [-0.11, 0.19]	<b>+0.23<sup>^</sup></b> [0.04, 0.45]	-0.02 [-0.18, 0.11]	<b>-0.26<sup>^</sup></b> [-0.43, -0.07]
Conscientiousness	+0.03 [-0.09, 0.20]	-0.14 [-0.33, 0.01]	-0.06 [-0.21, 0.09]	<b>+0.19<sup>^</sup></b> [0.01, 0.39]	+0.03 [-0.15, 0.16]	-0.05 [-0.19, 0.11]
Neuroticism	+0.11 [-0.01, 0.30]	+0.03 [-0.15, 0.22]	-0.02 [-0.17, 0.14]	-0.08 [-0.30, 0.10]	-0.09 [-0.29, 0.01]	+0.05 [-0.09, 0.23]
Agreeableness	+0.04 [-0.11, 0.20]	+0.01 [-0.18, 0.24]	+0.02 [-0.13, 0.17]	-0.03 [-0.27, 0.16]	-0.06 [-0.21, 0.10]	+0.03 [-0.18, 0.23]
Openness	0.00 [-0.16, 0.14]	+0.01 [-0.18, 0.23]	-0.07 [-0.24, 0.06]	0.00 [-0.22, 0.19]	+0.07 [-0.05, 0.25]	-0.01 [-0.22, 0.17]

*Note:*

Square brackets indicate 95% confidence intervals, estimated using bootstrapping with 10,000 samples. Coefficients with confidence intervals that do not include zero are marked with a caret (^) and are bolded.

**Table 3***Sample characteristics and materials for each of the three studies.*

Characteristics	Study		
	1	2	3
Time	Before peak bushfire severity	After peak bushfire severity	After peak bushfire severity
<b>Data collection dates</b>			
Start	24-Sep-2019	25-Feb-2020	13-Mar-2020
End	09-Nov-2019	02-Mar-2020	26-Mar-2020
<b>Sample characteristics</b>			
<i>n</i>	435	413	213
Mean age in years ( <i>SD</i> )	46.71 (17.77)	46.82 (18.04)	47.13 (17.29)
Number of women in sample (%)	219 (50.34%)	198 (47.94%)	103 (48.36%)
<b>Materials</b>			
Q-sort task	✓	✓	✓
Auxiliary psychological scales	✓	✗	✓
Fire Perception Scale	✗	✗	✓
Change in policy items	✗	✗	✓

**Table 4**

*Difference in means of climate change cognition and affect characteristics between Study 1 and Study 3.*

Psychological characteristics	$M_{Study\ 3} - M_{Study\ 1}$				
	Estimate	95% CI	t	p	$p_{adjusted}$
Perceptions of Natural Causes	0.39	[0.13, 0.65]	2.95	.003	.04*
Response Scepticism	0.19	[0.03, 0.35]	2.29	.022	.27
Perceived Effectiveness of Green Policies	-0.20	[-0.45, 0.05]	-1.60	.110	1.00
Worry about Climate Change	-0.11	[-0.28, 0.05]	-1.35	.178	1.00
Perceptions of Carbon Emission Causes	-0.15	[-0.38, 0.08]	-1.29	.197	1.00
Perceived Human Contribution	-0.18	[-0.46, 0.11]	-1.22	.222	1.00
Epistemic Scepticism	0.09	[-0.08, 0.25]	1.04	.300	1.00
Knowledge Volume	0.06	[-0.06, 0.19]	0.99	.325	1.00
Perceived Personal Consequences	0.12	[-0.12, 0.36]	0.97	.331	1.00
Perceptions of Environmental Harm Causes	-0.10	[-0.35, 0.16]	-0.75	.457	1.00
Perceived Effectiveness of Engineering Policies	-0.04	[-0.22, 0.14]	-0.43	.670	1.00
Perceived Effectiveness of Carbon Policies	-0.04	[-0.25, 0.18]	-0.33	.742	1.00
Perceived Societal Consequences	-0.01	[-0.25, 0.22]	-0.11	.914	1.00

*Note:*

\*  $p_{adjusted} < .05$ ;

*p* values were adjusted using the Holm (1979) method.

**Table 5**

*Items of the Fire Perception Scale, their loadings onto each factor, their mean score, and their standard deviation.*

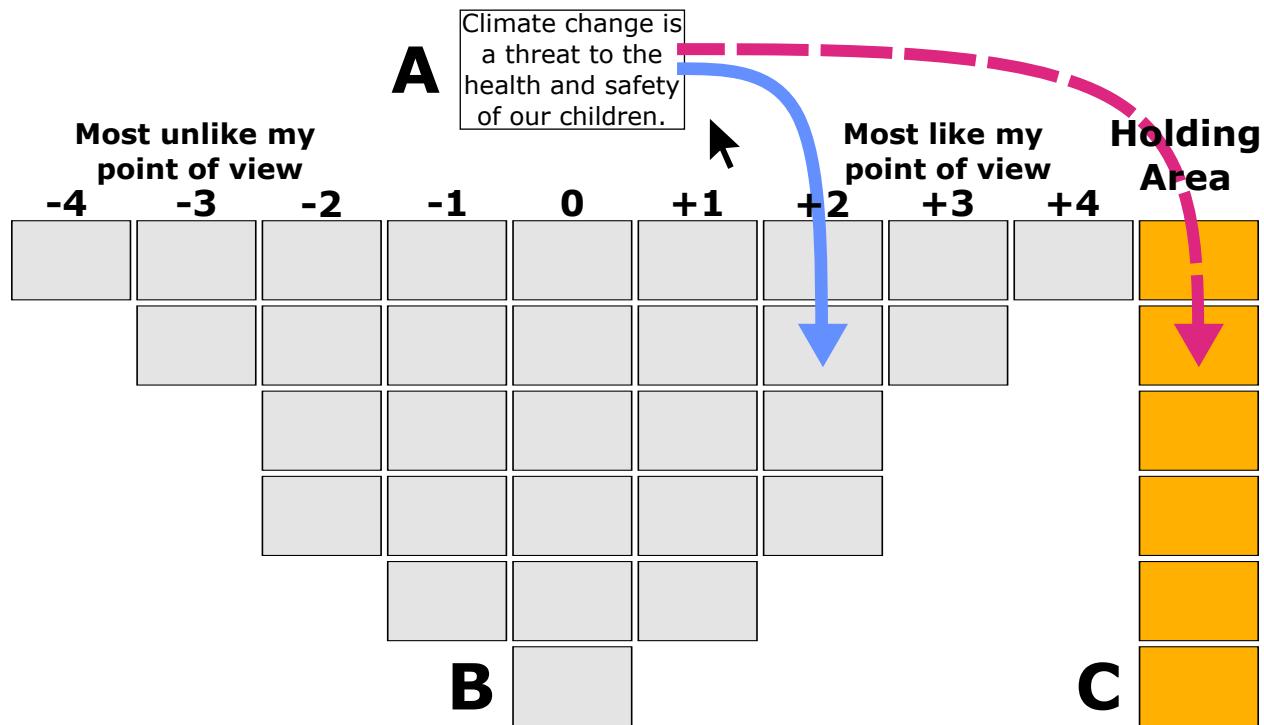
Item	Factors			Descriptives	
	Climate Processes	Fire Realities	Arson Causes	<i>M</i>	<i>SD</i>
1. Climate change made the 2019-20 Australian bushfires more severe	<b>0.78</b>	0.34	-0.22	3.62	1.40
2. Climate change made the 2019-20 Australian bushfires less likely to occur	0.27	<b>-0.70</b>	<b>0.42</b>	2.19	1.27
3. The 2019-20 Australian bushfires have accelerated climate change	<b>0.84</b>	0.05	-0.14	3.16	1.30
4. The 2019-20 Australian bushfires are severe	0.17	<b>0.86</b>	0.23	4.50	0.79
5. If the government increased taxes on all fossil fuels (e.g., gasoline, oil, coal, kerosene), Australia would be less likely to experience extreme bushfires	<b>0.84</b>	-0.19	0.13	2.55	1.32
6. If we changed our lifestyles to reduce our consumption, Australia would be less likely to experience bushfires	<b>0.86</b>	-0.06	0.08	3.05	1.39
7. Over one hundred arsonists have contributed to the 2019-20 Australian bushfires	-0.10	0.04	<b>0.94</b>	3.47	1.20

*Note:*

Bolded loadings are greater than .40 in magnitude.

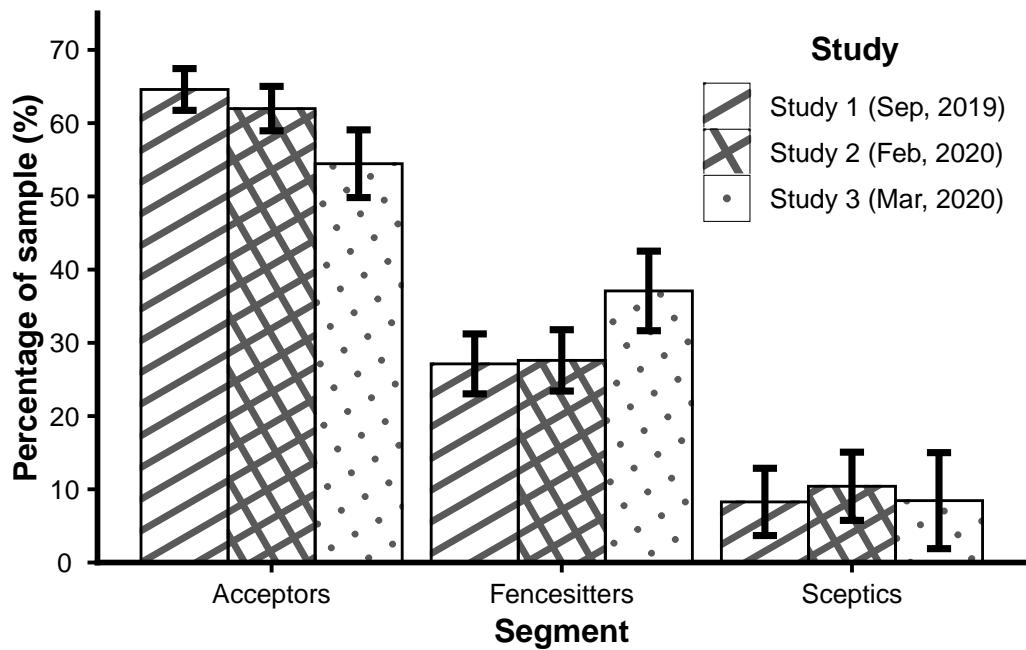
Figure 1

*Schematic of the Q-sort task. Participants read through a stack of statements (A) by dragging the top-most statement into the grey box that best corresponded to their point of view (B). As the majority of statements had to be placed around the midpoint, participants could only highlight a few statements that strongly reflect their point of view. Participants could rearrange statements at any time during the task. To facilitate this process, participants could temporarily place statements in the yellow holding area (C). Figure reproduced without changes from Andreotta et al. (2022), under the Creative Commons license (CC BY 4.0).*



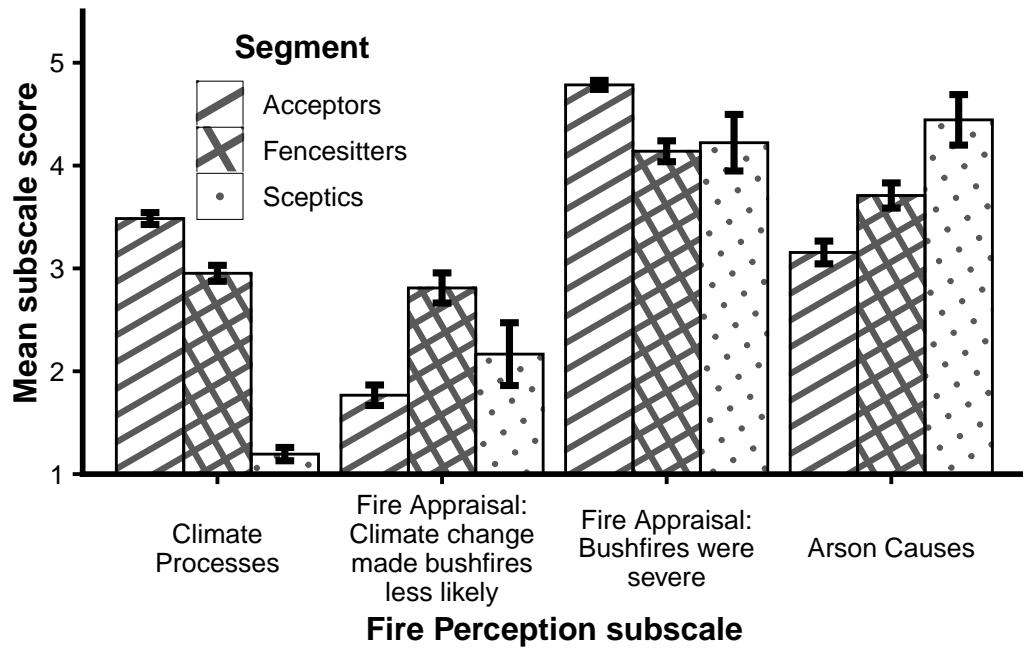
**Figure 2**

*The segment membership of each study, as a proportion (percentage) of the sample. Error bars indicate one standard error of the proportion.*



**Figure 3**

*Mean Fire Perception subscale scores as a function of segment. Error bars indicate one standard error above and below the mean.*



**Figure 4**

*Policy direction preferences as a proportion (percentage) of each segment. Error bars indicate one standard error of the proportion.*

